

UDC 519.21

DOI: <https://doi.org/10.17721/1812-5409.2025/2.10>

Dmytro TYKHONENKO, PhD Student

ORCID ID: 0009-0002-1144-6852

e-mail: tykhonenko.dmytro@gmail.com

Taras Shevchenko National University of Kyiv, Kyiv, Ukraine

Rostyslav YAMNENKO, DSc (Phys. & Math.)

ORCID ID: 0000-0002-9612-7959

e-mail: rostyslav.yamnenko@knu.ua

Taras Shevchenko National University of Kyiv, Kyiv, Ukraine

Katalin KUCHINKA, PhD (Phys. & Math.)

ORCID ID: 0000-0001-8129-9937

e-mail: kucsinka.katalin@unithe.hu

Ferenc Rakoczi II Transcarpathian Hungarian University, Berehove, Ukraine; University Tokaj, Sárospatak, Hungary

IMPROVEMENT OF CHARACTERIZATION INEQUALITY FOR NORMS OF φ -SUB-GAUSSIAN RANDOM VARIABLE

The paper is focused on refining the characterization inequality for the norm of φ -sub-Gaussian random variables. The core of this improvement lies in a detailed analysis of the remainder terms in Stirling's formula, which provides an approximation for $n!$. This study conducts a comparative analysis of several pairs of known remainder terms, α_n and β_n , to identify the pair that minimizes the approximation error. By calculating the absolute difference for these pairs, we demonstrate that the tightest bounds are achieved with a specific selection of higher-order terms. This optimized approximation for $n!$ is then applied to sharpen the constant in a characterization inequality for comparing different norms of a φ -sub-Gaussian random variable and also in the theorem on φ -sub-Gaussianity of a random variable.

Key words: Stirling's formula, remainder, characterization inequality, norm, φ -sub-Gaussian random variable.

AMS 2020 classification: 60E15, 26D15.

Introduction

Stirling's formula is a cornerstone of mathematical analysis, offering a powerful approximation for the factorial function, $n!$. The classical representation of this formula, known since the 18th century, states that as n approaches infinity, the ratio of $n!$ to $\sqrt{2\pi n}(n/e)^n$ converges to 1. This approximation is indispensable in various fields, including probability theory, statistics, and physics, where direct computation of factorials for large n is impractical.

While the limit-based formula is powerful, its practical application often requires more precise error bounds. Consequently, a significant body of research has been dedicated to refining Stirling's original approximation. The primary objective was to conduct a comparative analysis of several of these well-established remainder terms. The goal is to identify the specific pair of α_n and β_n that yields the least approximation error among the considered options.

The refined result is then applied to sharpen the characterization inequality for the norm of φ -sub-Gaussian random variables. These variables are a generalization of Gaussian random variables and have been widely researched and described in monographs (Buldygin, & Kozachenko, 2000; Vasylyk, Kozachenko, & Yamnenko, 2008).

1. Stirling's Formula Remainders

Since Stirling's time (1764), it has been known that $n! \sim C\sqrt{n}(n/e)^n$, with $C = \sqrt{2\pi}$, that is

$$\lim_{n \rightarrow \infty} \frac{n!}{\sqrt{2\pi n}(n/e)^n} = 1. \quad (1)$$

Formula (1) is known as Stirling's formula, and many elementary proofs of it have been given. Most such proofs use Wallis' formula

$$\lim_{n \rightarrow \infty} \frac{2 \cdot 2 \cdot 4 \cdot 4 \cdot \dots \cdot 2n \cdot 2n}{1 \cdot 3 \cdot 3 \cdot 5 \cdot \dots \cdot (2n-1)(2n+1)} = \frac{\pi}{2}, \quad (2)$$

to show that $C = \sqrt{2\pi}$. However, one of the elementary proofs by W. Feller (Feller, 1967) avoids any appeal to Wallis' formula. Since completely elementary proofs of (2) are well known (see, for example, (Schwartz, 1967)), this does not seem to be essential. In addition, a number of upper and lower bounds for $n!$, most of which imply (1), have been obtained by various authors. One of the most elementary estimates $e^{11/12}\sqrt{n}(n/e)^n < n! < e\sqrt{n}(n/e)^n$ was obtained by Hummel (Hummel, 1940) in 1940. Most bounds are of the form

$$\sqrt{2\pi n}(n/e)^n e^{\alpha_n} < n! < \sqrt{2\pi n}(n/e)^n e^{\beta_n}, \quad (3)$$

where α_n and β_n tend to zero through positive values. For example, $\beta_n = (12n)^{-1}$ was proved in (Uspensky, 1937), while the successively better values

$$\alpha_n^{(1)} = \frac{1}{\left(12n + \frac{3}{2(2n+1)}\right)}, \quad \alpha_n^{(2)} = \frac{1}{12n} - \frac{1}{360n^3}, \quad \alpha_n^{(3)} = \frac{1}{12n} - \frac{1}{360n^3} + \frac{1}{\left(1260\left(1 + \frac{2}{n(n+1)}\right)n^5\right)}, \quad (4)$$

$$\begin{aligned}\beta_n^{(1)} &= \frac{1}{12n} - \frac{1}{((360 + \gamma_n)n^3)}, \quad \text{where } \gamma_n = 30 \frac{7n(n+1) + 1}{n^2(n+1)^2}, \\ \beta_n^{(2)} &= \frac{1}{12n} - \frac{1}{(360n(n+1)(n+2))}, \\ \beta_n^{(3)} &= \frac{1}{12n} - \frac{1}{360n^3} + \frac{1}{1260n^5}\end{aligned}\quad (5)$$

were obtained in (Maria, 1965; Nanjundiah, 1965; Beesack, 1969; Hu, Lui, & Shi, 2006; Long, 2014; Macyz, 2021). As proofs for the mentioned remainders are given – we focus on choosing the left and right remainders that will give us the least $n!$ approximation error among those listed.

To do this, we need to calculate the possible error for each pair of left and right remainders. In common sense, it looks like $\frac{e^{\alpha n} - e^{\beta n}}{e^{\beta n}} = e^{\alpha n - \beta n} - 1$. And for an equation like this – we can use the fact that for small x , $e^x - 1 \approx x$.

Let's calculate the first pair of α and β with γ_n defined in (5).

$$\beta_n^{(1)} = \frac{1}{12n} - \frac{1}{(360 + \gamma_n)n^3} = \frac{1}{12n} - \frac{1}{(360 + \frac{210n^2 + 210n + 30}{(n^2 + n)^2})n^3} = \frac{33n^3 + 91n^2 + 123n + 60}{60n(7n^3 + 19n^2 + 25n + 12)}.\quad (6)$$

$$\alpha_n^{(1)} = \frac{1}{\left(12n + \frac{3}{2(2n+1)}\right)} = \frac{4n + 2}{48n^2 + 24n + 3}.\quad (7)$$

With the equations calculated above (6) and (7), we can compute the absolute difference between α and β .

$$\begin{aligned}|\alpha_n^{(1)} - \beta_n^{(1)}| &= \left| \frac{4n + 2}{48n^2 + 24n + 3} - \left(\frac{1}{12n} - \frac{1}{(360 + \gamma_n)n^3} \right) \right| \\ &= \left| \frac{32n^5 + 80n^4 + 31n^3 - 75n^2 - 123n - 60}{60n(112n^5 + 360n^4 + 559n^3 + 411n^2 + 121n + 12)} \right|.\end{aligned}\quad (8)$$

Now let's do the same for the next two pairs of α and β from (4) and (5) respectively.

$$|\alpha_n^{(2)} - \beta_n^{(2)}| = \left| \frac{1}{12n} - \frac{1}{360n^3} - \left(\frac{1}{12n} - \frac{1}{360n(n+1)(n+2)} \right) \right| = \left| \frac{n^2 - (n+1)(n+2)}{360n^3(n+1)(n+2)} \right| = \frac{3n + 2}{360n^3(n+1)(n+2)},\quad (9)$$

$$|\alpha_n^{(3)} - \beta_n^{(3)}| = \left| \frac{1}{1260 \left(\frac{n^2 + n + 2}{n(n+1)} \right) n^5} - \frac{1}{1260n^5} \right| = \left| \frac{n^2 + n - (n^2 + n + 2)}{1260(n^2 + n + 2)n^5} \right| = \frac{2}{1260(n^2 + n + 2)n^5}.\quad (10)$$

As we consider $n \geq 2$, it's easy to see that α and β achieve their supremum with $n = 2$. So, to calculate the least possible error – let's substitute it into equations (8), (9), (10):

$$\begin{aligned}|\alpha_n^{(1)} - \beta_n^{(1)}| &= \frac{1946}{120 \cdot 15714} \approx 1.03e^{-3}, \\ |\alpha_n^{(2)} - \beta_n^{(2)}| &= \frac{3 \cdot 2 + 2}{360 \cdot 8 \cdot 3 \cdot 4} \approx 2.3e^{-4}, \\ |\alpha_n^{(3)} - \beta_n^{(3)}| &= \frac{2}{1260 \cdot (4 + 2 + 2) \cdot 32} \approx 6.2e^{-6}.\end{aligned}\quad (11)$$

As we convinced earlier, $\alpha_n^{(3)}$ and $\beta_n^{(3)}$ from (4) and (5) respectively give us the least error (11). That is why we will be using it further. Now let's assume that $n = 2$.

Then

$$\beta_n^{(3)} = \frac{1}{12 \cdot 2} - \frac{1}{360 \cdot 8} + \frac{1}{1260 \cdot 32} = \frac{1680 - 14 + 1}{40320} = \frac{1667}{40320}$$

and

$$\alpha_n^{(3)} = \frac{1}{12 \cdot 2} - \frac{1}{360 \cdot 8} + \frac{2 \cdot 3}{1260 \cdot (2 \cdot 3 + 2) \cdot 32} = \frac{13334}{322560} = \frac{6667}{161280}.\quad (12)$$

2. Improvements of inequality for norm of space of φ -sub-Gaussian random variables

φ -sub-Gaussian random variables, spaces, and their properties were introduced in (Giuliano-Antonini, Kozachenko, & Nikitina, 2003). We will try to improve the accuracy of the inequality for the norms of the space of φ -sub-Gaussian random variables, referring to the lemmas from there.

Lemma 1. Let $\xi \in \text{Sub}_\varphi(\Omega)$. Then for every $a > 0$ the next inequality holds true:

$$\mathbb{E} |\xi|^a \leq 2 \left(\frac{a}{e} \right)^a (\tau_\varphi(\xi))^a \inf_{t>0} \exp(\varphi(t - a \ln t)) \leq 2 (\tau_\varphi(\xi))^a \left(\frac{a}{\varphi^{-1}(a)} \right)^a.$$

Then, from 1, we can obtain the following result.

Corollary 1. Let $\xi \in Sub_{\varphi}(\Omega)$. Then the next inequality holds true

$$\tau_{\varphi}(\xi) \geq \theta_{\varphi}(\xi)/\sqrt{2},$$

where $\theta_{\varphi}(\xi) = \sup_{n \geq 2} E(|\xi|^n)^{\frac{1}{n}} \frac{\varphi^{-1}(n)}{n}$, and $\theta_{\varphi}(\xi)$ is a norm for $Sub_{\varphi}(\Omega)$.

Lemma 2. Let $\xi \in Sub_{\varphi}(\Omega)$. Then for $k = 1, 2, \dots$ the following inequalities hold true:

$$|E \xi^k| \leq E |\xi|^k \leq 2 (\tau_{\varphi}(\xi))^k \frac{e^k}{(\varphi^{-1}(k))^k} k!.$$

Corollary 2. If $\xi \in Sub_{\varphi}(\Omega)$, then the next inequality holds true

$$\tau_{\varphi}(\xi) \geq \frac{1}{e\sqrt{2}} \nu_{\varphi}(\xi),$$

where $\nu_{\varphi}(\xi) = \sup_{n \geq 2} |E \xi^n|^{\frac{1}{n}} \frac{\varphi^{-1}(n)}{(n!)^{\frac{1}{n}}}$.

Norms $\theta_{\varphi}(\cdot)$ and $\nu_{\varphi}(\cdot)$ are less popular than $\tau_{\varphi}(\cdot)$, yet they are used in applications of φ -sub-Gaussian processes, see, for example (Troszki, 2017).

Lemma 3. Let $\xi \in Sub_{\varphi}(\Omega)$ and $\nu_{\varphi}(\xi) \geq \theta_{\varphi}(\xi)$. Then the next inequality holds true:

$$\nu_{\varphi}(\xi) \leq e^{\frac{329227}{322560}} \theta_{\varphi}(\xi). \quad (13)$$

Proof. We can see that this inequality holds true according to proposition 1, lemma 3, and Stirling's formula (3), stating that $n! = n^n e^{-n} (2\pi n)^{1/2} e^{\theta_n}$ for $n \geq 2$, where $|\theta_n| \leq \frac{1}{12n} - \frac{1}{360n^3} + \frac{1}{(1260(1 + \frac{2}{n(n+1)})n^5)}$. Using the value calculated in (12), we derive the following inequality

$$(n!)^{-\frac{1}{n}} = \frac{1}{n} \frac{e^{1+\frac{\theta_n}{n}}}{(2\pi n)^{\frac{1}{2n}}} \leq \frac{1}{n} e^{1+\frac{\theta_n}{n}} = \frac{1}{n} e^{\frac{329227}{322560}}.$$

□

Lemma 3 improves the following result from (Giuliano-Antonini, Kozachenko, & Nikitina, 2003, Corollary 4.4):

$$\nu_{\varphi}(\xi) \leq \exp \left\{ \frac{49}{48} \right\} \theta_{\varphi}(\xi).$$

Lemma 4. Let ξ be a random variable with $E \xi = 0$, and let φ be a N -function satisfying the condition Q . Let $\lambda_0 > 0$ be some number and

$$c_0 = \inf_{0 < |\lambda| \leq \lambda_0} \varphi(\lambda)/\lambda^2.$$

Let also $\nu_{\varphi}(\xi) < \infty$. Then $\xi \in Sub_{\varphi}(\Omega)$ and the following inequality holds:

$$\tau_{\varphi}(\xi) \leq S_{\varphi} \nu_{\varphi}(\xi), \quad (14)$$

where $S_{\varphi} = \max_{i=1,3} \gamma_i^{-1}$, and γ_1 is the root of the equation $\gamma = \lambda_0 \sqrt{c_0(1-\gamma)}$, γ_2 is the root of the equation $\gamma^3 - 2(1-\gamma) = 0$, and γ_3 is the root of the equation $\gamma = \varphi^{-1}(2) \sqrt{c_0(1-\gamma)}$.

The proof can be found in (Giuliano-Antonini, Kozachenko, & Nikitina, 2003).

Also, note that Lemmas 3 and 4 establish the equivalence of the norms $\nu_{\varphi}(\xi)$ and $\tau_{\varphi}(\xi)$. Furthermore, a random variable ξ belongs to $Sub_{\varphi}(\Omega)$ if and only if $E \xi = 0$ and $\nu_{\varphi}(\xi) < \infty$.

3. Orlicz spaces of exponential type

Let ψ be an arbitrary N -function. The Orlicz space generated by the N -function $U(x) = \{\exp \{\psi(x)\} - 1, x \in \mathbb{R}\}$, is called an Orlicz space of exponential type. We shall be interested in the Orlicz space of exponential type $Exp_{\varphi^*}(\Omega)$ generated by the Young-Fenchel transform φ^* of an N -function φ . The Luxemburg norm in $Exp_{\varphi^*}(\Omega)$ is denoted by σ_{φ^*} , i.e. for a random variable $\xi \in Exp_{\varphi^*}(\Omega)$ we have

$$\sigma_{\varphi^*}(\xi) = \inf \{a > 0: E \{\exp \varphi^*(\xi/a)\} \leq 2\}.$$

The space $Exp_{\varphi^*}(\Omega)$ is a Banach space with respect to the norm $\sigma_{\varphi^*}(\cdot)$.

Lemma 5. Let φ be an N -function and $\xi \in Exp_{\varphi^*}(\Omega)$. Then, for any $p \geq 1$, the following inequality holds:

$$(E |\xi|^p)^{\frac{1}{p}} \leq 2^{\frac{1}{p}} \frac{p}{\varphi^{-1}(p)} \sigma_{\varphi^*}(\xi). \quad (15)$$

The proof can be found in (Giuliano-Antonini, Kozachenko, & Nikitina, 2003).

As a consequence of (13), the following inequalities are also improved.

Lemma 6. Let φ be a N -function satisfying the condition Q , and let ξ be a random variable such that $\xi \in \text{Sub}_\varphi(\Omega)$ and $E\xi = 0$. Then

$$\tau_\varphi(\xi) \leq S_\varphi e^{\frac{329227}{322560}} \sigma_{\varphi^*}(\xi),$$

where S_φ is defined in (14).

Lemma 7. Let ξ be a random variable such that

$$P\{|\xi| \geq x\} \leq C \exp\left\{-\psi\left(\frac{x}{D}\right)\right\},$$

where $\psi(x)$ is an N -function. Then $\xi \in \text{Exp}_\psi(\Omega)$ and $\sigma_\psi(\xi) \leq (1 + C)D$.

Lemma 8. Let $\xi \in \text{Sub}_\varphi(\Omega)$; then $\xi \in \text{Exp}_{\varphi^*}(\Omega)$ and the following inequality holds:

$$\sigma_{\psi^*}(\xi) \leq 3\tau_\varphi(\xi). \quad (16)$$

Considering the lemmas, propositions, and equation (12) mentioned above, we can arrive at the following theorem.

Theorem 1. A random variable ξ belongs to the space $\text{Sub}_\varphi(\Omega)$ only if $E\xi = 0$ and there exist constants $C > 0$ and $D > 0$ such that

$$P\{|\xi| \geq x\} \leq C \exp\left\{-\varphi^*\left(\frac{x}{D}\right)\right\} \quad (17)$$

for any $x > 0$. If (17) holds, then

$$\tau_\varphi(\xi) \leq S_\varphi e^{\frac{329227}{322560}} (1 + C)D,$$

where S_φ is defined in (14).

Example. Consider the Orlicz N -function $\varphi_p(x) = \frac{|x|^p}{p}$, $1 < p < 2$, or, for $p > 2$, the piecewise definition $\varphi_p(x) = \frac{|x|^p}{p}$, $|x| \geq 1$, and $\varphi_p(x) = \frac{|x|^2}{p}$, $|x| < 1$, and examine inequality (16) in this particular setting.

Put $\tau_{\varphi_p}(\xi) = \tau$. Let's apply Lemma 1. Since for $s > 1$

$$\inf_{t \geq 1} \exp\{\varphi_p(t) - s \ln t\} = \left(\frac{e}{s}\right)^{\frac{s}{p}},$$

then

$$E|\xi|^s \leq 2 \left(\frac{s}{e}\right)^{\frac{s}{q}} \tau^s. \quad (18)$$

Since $\varphi_p^*(x) = \frac{|x|^q}{q}$ for the case $1 < p < 2$, and $\varphi_p^*(x) \leq \frac{|x|^q}{q}$ (as $q < 2$) we have by (18) for $a > 0$

$$E \exp\{\varphi^*(\xi/a)\} \leq E \exp\left\{\frac{|\xi|^q}{qa^q}\right\} = \sum_{k=0}^{\infty} \frac{1}{k!} \frac{E|\xi|^k}{(qa^q)^k} \leq 1 + 2 \sum_{k=1}^{\infty} \left(\frac{k^k}{k!e^k}\right) \left(\frac{\tau}{a}\right)^q k. \quad (19)$$

It follows from Stirling's formula and (12) that

$$\frac{k^k}{k!e^k} \leq \frac{1}{\sqrt{2\pi}} e^{\frac{13334}{322560}}.$$

Therefore

$$E \exp\left\{\varphi^*\left(\frac{\xi}{a}\right)\right\} \leq 1 + \frac{2e^{\frac{13334}{322560}}}{\sqrt{2\pi}} \sum_{k=1}^{\infty} \left(\frac{\tau}{a}\right)^{kq}.$$

The above geometric series converges if $\tau/a < 1$ and in such case its sum is equal to

$$1 + \frac{2e^{\frac{13334}{322560}}}{\sqrt{2\pi}} \frac{(\tau/a)^q}{1 - (\tau/a)^q}.$$

Put

$$L = \left(\frac{2e^{\frac{13334}{322560}}}{\sqrt{2\pi}} + 1\right)^{1/q},$$

then the last quantity in (19) is not greater than 2 if $a/\tau > L > 1$. This gives us the following result

$$\sigma_{\varphi^*}(\xi) \leq L\tau_{\varphi_p}(\xi).$$

Discussions and conclusion

By refining Stirling's formula through an optimal choice of remainder terms, we obtained sharper constants in the characterization inequality for norms of φ -sub-Gaussian random variables. This illustrates how classical analytic improvements translate into tighter probabilistic bounds, with potential applications to concentration inequalities and stochastic modeling.

Authors' contribution: Rostyslav Yamnenko — conceptualization and methodology, Dmytro Tykhonenko — analysis of sources, theoretical foundations of research, writing an original draft, Katalin Kuchinka — analysis of results, viewing and editing.

Sources of funding. Funding is partially provided by Taras Shevchenko National University of Kyiv.

References

- Beesack, P. (1969). Improvements of Stirling's formula by elementary methods. *Publikacije Elektrotehničkog fakulteta. Serija Matematika i fizika*, 274/301, 17–21.
- Buldygin, V. V., & Kozachenko, Y. V. (2000). *Metric characterization of random variables and random processes*. American Mathematical Society, Providence, RI. <https://doi.org/10.1090/mmono/188>
- Feller, W. (1967). A direct proof of Stirling's formula. *American Mathematical Monthly*, 74(10), 1223–1225. <https://doi.org/10.2307/2315671>
- Giuliano-Antonini, R., Kozachenko, Y. V., & Nikitina, T. (2003). Space of φ -sub-Gaussian random variables. *Rendiconti, Accademia Nazionale delle Scienze detta dei XL, Memorie di Matematica e Applicazioni*, XXVII(121°), 92–124.
- Hu, M., Lui, F., & Shi, X. (2006). A new asymptotic series for the Gamma function. *Journal of Computational and Applied Mathematics*, 195(1), 134–154. <https://doi.org/10.1016/j.cam.2005.03.081>
- Hummel, P. (1940). A note on Stirling's formula. *American Mathematical Monthly*, 47(2), 97–99. <https://doi.org/10.2307/2303360>
- Long, L. (2014). On Stirling's formula remainder. *Applied Mathematics and Computation*, 247, 494–500. <https://doi.org/10.1016/j.amc.2014.09.012>
- Macyz, J. (2021). On Stirling's formula. *Lietuvos matematikos rinkinys*, 47, 526–530. <https://doi.org/10.15388/LMR.2007.24256>
- Maria, A. (1965). A remark on Stirling's formula. *American Mathematical Monthly*, 72(10), 1096–1098. <https://doi.org/10.2307/2315957>
- Nanjundiah, T. (1965). Note on Stirling's formula. *American Mathematical Monthly*, 66(8), 701–703. <https://doi.org/10.2307/2309346>
- Schwartz, A. (1967). *Calculus and analytic geometry* (2nd ed.). Holt, Rinehart & Winston.
- Troshki, N. (2017). Upper bounds for supremums of the norms of the deviation between a homogeneous isotropic random field and its model. *Theory of Probability and Mathematical Statistics*, 94, 159–184. <https://doi.org/10.2307/2309346>
- Uspensky, J. (1937). *Introduction to mathematical probability*. McGraw Hill.
- Vasylyk, O., Kozachenko, Y., & Yamnenko, R. (2008). φ -sub-Gaussian random processes. VPC "Kyivskyi Universytet" [in Ukrainian]. [Василик, О., Козаченко, Ю., & Ямненко, Р. (2008). φ -субгауссові випадкові процеси. ВПЦ "Київський університет"].

Отримано редакцією журналу / Received: 10.07.25
Прорецензовано / Revised: 12.09.25
Схвалено до друку / Accepted: 10.10.25

Дмитро ТИХОНЕНКО, асп.
ORCID ID: 0009-0002-1144-6852
e-mail: tykhonenko.dmytro@gmail.com
Київський національний університет імені Тараса Шевченка, Київ, Україна

Ростислав ЯМНЕНКО, д-р фіз.-мат. наук
ORCID ID: 0000-0002-9612-7959
e-mail: rostyslav.yamnenko@knu.ua
Київський національний університет імені Тараса Шевченка, Київ, Україна

Каталін КУЧІНКА, канд. фіз.-мат. наук
ORCID ID: 0000-0001-8129-9937
e-mail: kucsinka.katalin@unithe.hu
Закарпатський угорський університет імені Ференца Ракоці II, Берегове, Україна; Токайський університет, Шарошпатак, Угорщина

УТОЧНЕННЯ ХАРАКТЕРИЗАЦІЙНОЇ НЕРІВНОСТІ ДЛЯ НОРМ φ -СУБГАУССОВОЇ ВИПАДКОВОЇ ВЕЛИЧИНИ

Здійснено уточнення характеристичної нерівності між нормами φ -субгауссової випадкової величини. Суть цього вдосконалення полягає в більш детальному аналізі залишкових членів у формулі Стірлінга, що використовується для підрахунку наближеного значення для $n!$. У пропонуваному дослідженні проведено порівняльний аналіз декількох пар відомих залишкових членів, α_n і β_n , щоб визначити пару, яка мінімізує похибку наближення. Обчислюючи абсолютну різницю для цих пар, продемонстровано, що найтісніші межі досягаються у разі конкретного вибору членів вищого порядку. Це поліпшене наближення для $n!$ потім застосовується для уточнення сталої в характеристичній нерівності для порівняння різних норм φ -субгауссової випадкової величини та теоремі про φ -субгауссовість випадкової величини.

Ключові слова: формула Стірлінга, залишок, характеристична нерівність, норма, φ -субгауссова випадкова величина.

Автори заявляють про відсутність конфлікту інтересів. Спонсори не брали участі в розробленні дослідження (у зборі, аналізі чи інтерпретації даних, якщо це мало місце), у написанні рукопису та в рішенні про публікацію результатів.

The authors declare no conflicts of interest. The funders had no role in the design of the study (in the collection, analyses or interpretation of data if applicable), in the writing of the manuscript as well as in the decision to publish the results.